Large Scale Hypothesis Testing

CSE545 - Spring 2022 Stony Brook University

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Big Data Analytics, The Class

Goal: Generalizations A model or summarization of the data.

Data Workflow Systems

Algorithms and Analyses

Hadoop File System S Streaming Spark

MapReduce

Tensorflow

Similarity Search

Link Analysis Large Scale Hyp. Testing

Recommendation Systems

Deep Learning

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The Data Whisperer

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Hypotheses! Potential findings -- to be tested for happenstance.

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Also known as... "Don't be Dilbert's Boss!"

Hypothesis -- something one asserts to be true.

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Formally, two types:

H₀: null hypothesis -- some "default" value; "null": nothing changes

 H_1 : the alternative -- the opposite of the null => a change or difference

H_o: null hypothesis -- some "default" value; "null": nothing changes

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Given null, what is the probability of the observation or worse

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Output: decision

Input: H_{α} , obs, α

observations (i.e. data) level of significance



```
Input: H_0, obs, \alpha
```

```
if p(x>=obs | H<sub>0</sub>) < α:
    decision = "Reject H<sub>0</sub>!"
else:
    decision = "Accept H<sub>0</sub>."
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Input: H_0 , obs, α

Conditional is sometimes evaluated indirectly by first finding the "critical value" of some <u>measurement</u> such that: if measurement > critical_value then p(obs/H0) < **Q**

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Input: H₀, obs, α

Need to estimate -

What is the distribution of values we would expect if the null was true? -- the "null distribution"

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X: A mapping from Ω to \mathbb{R} that describes the question we care about in practice.

X is a *continuous random variable* if it can take on an infinite number of values between any two given values. X is a *discrete random variable* if it takes only a countable number of values.

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Error of RedImageGenNet Classifier

Amount of sales of a blue case

X is a *continuous random variable* if it can take on an infinite number of values between any two given values.

X is a *continuous random variable* if there exists a function *fx* such that:

$$f_X(x) \ge 0$$
, for all $x \in X$,
 $\int_{-\infty}^{\infty} f_X(x) dx = 1$, and
 $P(a < X < b) = \int_a^b f_X(x) dx$

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fx : "probability density function" (pdf)



How to model?







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Common Trap

- $f_X(x)$ does not yield a probability • $\int_a^b f_X(x) dx$ does
 - x may be anything (\mathbb{R})
 - thus, $f_X(x)$ may be > 1



Common pdfs: Normal(0, 1)

 $P(-1 \le Z \le 1) \approx .68, \quad P(-2 \le Z \le 2) \approx .95, \quad P(-3 \le Z \le 3) \approx .99$



Credit: MIT Open Courseware: Probability and Statistics

Common pdfs: Normal(0, 1) ("standard normal")

How to "standardize" any normal distribution:

- 1. subtract the mean, μ (aka "mean centering")
- 2. divide by the standard deviation, $\boldsymbol{\sigma}$

 $z = (x - \mu) / \sigma$, (aka "z score")

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Discrete Random Variables

For a given *discrete* random variable X, *probability mass function (pmf)*, *fx:* $\mathbb{R} \rightarrow [0, 1]$, is defined by:

 $f_X(x) = \mathcal{P}(X = x)$

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Was a single sale a blue case: {0, 1}
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Was a single sale a blue case: {0, 1}

For a given random variable X, the *cumulative distribution function* (CDF), *Fx:* $\mathbb{R} \rightarrow [0, 1]$, is defined by:

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Discrete RVs

For a given random variable X, the cumulative distribution function (CDF), $Fx: \mathbb{R} \to [0, 1]$, is defined by: $F_X(x) = P(X \le x)$

For a given *discrete* random variable X, *probability mass function (pmf)*, *fx:* $\mathbb{R} \rightarrow [0, 1]$, is defined by:

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X is a *discrete random variable* if it takes only a countable number of values.

$$\sum_{i} f_X(x) = 1$$
$$F_X(x) = P(X \le x) = \sum_{x_i \le x} f_X(x)$$

Input: H_0 , observations, α

Need to estimate -

What is the distribution of values we would expect if the null was true? -- the "null distribution"

```
if p(x>=obs | H<sub>0</sub>) < a:
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H: The blue case is not selling more than average.



H_o: The blue case is not selling more than average. 50 sales; 2 colors (blue and red); Thus, average would be 25 blue sales

Input: H_0 , obs, α

null_dist = distribution of expected values under H_{ρ}

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if p(x>=obs | H<sub>0</sub>) < α:
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Output: decision
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H_o: The blue case is not selling more than average. **Observed 33 blue sales** 50 sales; 2 colors (blue and red); Thus, average would be 25 blue sales



```
Input: H_0, obs, \alpha
```

null_dist = distribution of expected values under H₀
p(x>=obs | H₀) =
if p(x>=obs | H₀) < α:
 decision = "Reject H₀!"
else:
 decision = "Accept H₀."
Output: decision

Input: H_0 , obs, α

null_dist = distribution of expected values under H_{ρ}

 $p(x \ge obs | H_{g}) = sum(pmf(null_dist, o) for o in range(obs,))$

```
if p(x>=obs | H<sub>0</sub>) < α:
    decision = "Reject H<sub>0</sub>!"
else:
    decision = "Accept H<sub>0</sub>."
Output: decision
```

Input: H_0 , obs, α

null_dist = distribution of expected values under H_{ρ}

 $p(x \ge obs | H_{\rho}) = 1 - cdf(null_dist, obs)$

```
if p(x>=obs | H<sub>0</sub>) < α:
    decision = "Reject H<sub>0</sub>!"
else:
    decision = "Accept H<sub>0</sub>."
Output: decision
```



Input: H_0 , obs, α

null_dist = distribution of expected values under H_{ρ}

 $p(x \ge 0.016 | H_0) = 1 - cdf(null_dist, obs) = 0.016$

if p(x>=obs | H_a) < α: decision = "Reject H_a!" else: decision = "Accept H_a." Output: decision





Input: H_{α} , obs, α null_dist = distribution of expected values under H_{a} $p(x \ge 0.239) = 1 - cdf(null_dist, obs) = 0.239$ if $p(x \ge obs | H_{\rho}) < \alpha$: 0.8 decision = "Reject H_{a} !" else: decision = "Accept H_a." 0.2 **Output:** decision 7 and N=20 0

```
Input: H_0, obs, \alpha
```

null_dist = distribution of expected values under H_{ρ}

```
p(x \le obs | H_{\rho}) = cdf(null_dist, obs)
```

```
if p(x<=obs | H<sub>0</sub>) < α:
    decision = "Reject H<sub>0</sub>!"
else:
    decision = "Accept H<sub>0</sub>."
Output: decision
```

```
Input: H_0, obs, \alpha
```

null_dist = distribution of expected obs under H₀
p(x<=obs | H₀) = cdf(null_dist, obs)
if p(x<=obs | H₀) < α:
 decision = "Reject H₀!"
else:
 decision = "Accept H₀."
Output: decision

```
Input: H_0, obs, \alpha
```

```
obs_ts = test_stat(obs)
null_dist = distribution of expected obs under H<sub>0</sub>
p(x<=obs | H<sub>0</sub>) = cdf(null_dist, obs)
if p(x<=obs | H<sub>0</sub>) < α:
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 $P(D|H_0)$: Given null, what is the probability of the observed data or worse? -> If <u>low enough</u>, then we "reject the null (H_0) in favor of H_1 ."



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A **p-value** (shaded green area) is the probability of an observed (or more extreme) result assuming that the null hypothesis is true.

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Set of possible results

A **p-value** (shaded green area) is the probability of an observed (or more extreme) result assuming that the null hypothesis is true.

(thanks, Wikipedia)

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A general framework for answering (yes/no) questions!

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- Are height and baldness related?
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A general framework for answering (yes/no) questions!

- Are height and baldness related?
- Is my deep predictive model better than the state of the art?
- Is the heat index of a community related to poverty?
- Is the heat index of a community related to poverty controlling for education rates?
- Does my website receive a higher average number of monthly visitors?

Failing to "reject the null" does not mean the null is true.

Why?

A general framework for answering (yes/maybe) questions!

- Are height and baldness related?
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- Is the heat index of a community related to poverty controlling for education rates?
- Does my website receive a higher average number of monthly visitors?

Failing to "reject the null" does not mean the null is true. However, if the sample is large enough, it may be enough to say that the effect size (correlation, difference value, etc...) is not very meaningful.

A general framework for answering (yes/maybe) questions!

• Are height and baldness related?

Why?

- Is my deep predictive model better than the state of the art?
- Is the heat index of a community related to poverty?
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General Question: Which fish do cats like?







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N = 50 cats







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 H_1 : Most cats like redfish. H_0 : Most cats don't like redfish.
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N = 50 cats; 32 like redfish; p = 0.016



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H: Most c

Now suppose instead of just redfish, you wanted to ask the same question for 10 kinds of fish: $H_{1,1}$: Most cats like redfish; $H_{1,2}$: Most cats like bluefish; $H_{1,3}$: Most cats like orangefish; ... with $\alpha = 0.05$, can you still conclude most cats like redfish?

don't like redfish.

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don't like redfish.

General Question: Which fish do cats like?

N = 50 cats; 32 like redfish; p = 0.016

 α = 0.05 -- probability threshold for happening upon the result even if it really doesn't exist.

What is the probability we happen upon once in ten times?

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-1 - p(not happening upon the result) = 1 - (1 - .05)¹⁰ = 1 - 0.599 = .4

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How to fix?

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Is there a way we could adjust alpha to keep it low enough?

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= 1 - 0.599 = .4

How to fix? 1- (1 - adjust(.05))¹⁰ < .05

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pon once in ten times? = $1 - (1 - .05)^{10}$ = 1 - 0.599

don't like redfish.

How to fix? 1 - (1 - (.05/10))¹⁰ = .0488

Type I, Type II Errors

| | | True state of nature | |
|----------|----------------|----------------------|------------------|
| | | H_0 | H_A |
| Our | Reject H_0 | Type I error | correct decision |
| decision | 'Accept' H_0 | correct decision | Type II error |
| | | | |

significance level ("p-value") = P(type I error) = P(Reject H₀ | H₀)
(probability we are incorrect)



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significance level ("p-value") = P(type I error) = P(Reject H₀ | H₀)
(probability we are incorrect)

power = 1 - P(type II error) = P(Reject H₀ | H₁)
(probability we are correct)

| | H_0 | H_A | |
|--------------|--|--|--|
| Reject H_0 | P(Reject H ₀ H ₀) | P(Reject H ₀ H ₄) | |

| | | True state of nature | |
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FWER: Family-wise error rate (Bonferroni Corrects) The probability of making >=1 type 1 error. $FWER = Pr(type1s>0) = 1 - Pr(type1s=0) = 1 - (1 - a)^m$



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 $1 - (1 - (.05/10))^{10} = .0488$



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FDR: False discovery rate (Benjamini-Hochberg corrects) type1s / (type1s + correctRejects)

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| | é. | 36 | |

FWER: Family-wise error rate (Bonferroni corrects) The probability of making >=1 type 1 error. $FWER = Pr(type1s>0) = 1 - Pr(type1s=0) = 1 - (1 - a)^m$

FDR: False discovery rate (Benjamini-Hochberg corrects) type1s / (type1s + correctRejects)

Proportion of false positives among *all* significant results.

The Hypothesis Test "Algorithm"

```
Input: H_0, obs, \alpha
```

```
obs_ts = test_stat(obs)
null_dist = distribution of expected test_stat under H<sub>0</sub>
p(x<=obs_ts | H<sub>0</sub>) = cdf(null_dist, obs_ts)
if p(x<=obs_ts | H<sub>0</sub>) < α:
    decision = "Reject H<sub>0</sub>!"
else:
    decision = "Accept H<sub>0</sub>."
Output: decision
```

The Multi-test "Algorithm"

```
Input: H_{\alpha}s, obs, \alpha
decisions = []
\alpha_{adj} = adjust(\alpha)
for H<sub>a</sub> in H<sub>a</sub>s
    obs ts = test_stat(obs)
    null dist = distribution of expected test_stat under H_{a}
    p(x<=obs_ts | H<sub>o</sub>) = cdf(null_dist, obs_ts)
    if p(x < = obs_ts | H_{\rho}) < \alpha_adj:
       decisions.append("Reject H<sub>a</sub>!")
    else:
       decisions.append("Accept H_{a}.")
Output: decisions
```

The Multi-test "Algorithm"

```
Input: H_{\alpha}s, obs, \alpha
decisions = []
\alpha_{adj} = adjust(\alpha) \#e.g. adjust(\alpha) = \alpha/len(H_{\alpha}s)
for H_{a} in H_{a}s
   obs_ts = test_stat(obs)
   null dist = distribution of expected test_stat under H_{a}
    p(x<=obs_ts | H<sub>o</sub>) = cdf(null_dist, obs_ts)
   if p(x <= obs_ts | H_{\rho}) < \alpha_adj:
       decisions.append("Reject H<sub>a</sub>!")
    else:
       decisions.append("Accept H_{a}.")
Output: decisions
```

Multi-test "Algorithm" Alternative

```
Input: H_{\alpha}s, obs, \alpha
decisions = []
for H_{a} in H_{a}s
   obs ts = test_stat(obs)
   null dist = distribution of expected test_stat under H_{a}
   p(x \le obs_ts | H_{\rho}) = cdf(null_dist, obs_ts)
   p_adj = inverse_adjust(p(x<=obs_ts|H<sub>a</sub>))#e.g. p*len(H<sub>a</sub>s)
   if p adj < \alpha:
      decisions.append("Reject H<sub>a</sub>!")
    else:
      decisions.append("Accept H_{a}.")
Output: decisions
```

Statistical Considerations for Big Data

- Average multiple models (ensemble techniques)
- 2. Correct for multiple tests (Bonferonni's Principle)
- 3. Smooth data
- 4. "Plot" data (or figure out a way to look at a lot of it "raw")
- 5. Interact with data

- 6. Know your "real" sample size
- 7. Correlation is not causation
- 8. Define metrics for success (set a baseline)
- 9. Share code and data
- 10. The problem should drive solution

(http://simplystatistics.org/2014/05/22/10-things-statistics-taught-us-about-big-data-analysis/)

Comparing Variables

• Linear Regression

- Pearson Product-Moment Correlation
- Multiple Linear Regression
- (Multiple) Logistic Regression
- Ridge Regression (L2 Penalized)
- Lasso Regression (L1 Penalized)

Comparing Variables

Finding a linear function based on X to best yield Y.

X = "covariate" = "feature" = "predictor" = "regressor" = "independent variable"

Y = "response variable" = "outcome" = "dependent variable"

Regression:
$$r(x) = E(Y|X = x)$$

goal: estimate function r

The **expected** value of *Y*, given that the random variable *X* is equal to some specific value, *x*.

Finding a linear function based on X to best yield Y.

X = "covariate" = "feature" = "predictor" = "regressor" = "independent variable"

Y

Y = "response variable" = "outcome" = "dependent variable"

Regression:
$$r(x) = E(Y|X = x)$$

goal: estimate the function r

Linear Regression (univariate version):

goal: find β_0 , β_1 such that

$$r(x) = \beta_0 + \beta_1 x$$

$$r(x) \approx \mathbf{E}(Y|X = x)$$

Linear Regression

Simple Linear Regression
$$Y_i = \beta_0 + \beta_1 X_i + \epsilon_i$$

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How to estimate intercept (\Box_0) and slope intercept (\Box_1) ?

Least Squares Estimate. Find $\hat{\beta}_0$ and $\hat{\beta}_1$ which minimizes the *residual sum of squares:*

$$J(\Box) = RSS = \sum_{i=1}^{n} \hat{\epsilon}_{i}^{2} = \sum_{i=1}^{n} (Y_{i} - \hat{Y}_{i})^{2} = \sum_{i=1}^{n} (Y_{i} - \hat{\beta}_{0} - \hat{\beta}_{1}X_{i})^{2}$$



the residual sum of squares:

$$\mathcal{I}(\Box) = RSS = \sum_{i=1}^{n} (Y_i - \hat{Y}_i)^2$$

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(http://rasbt.github.io/mlxtend/user guide/general concepts/gradient-optimization/)











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Linear Regression



Pearson Product-Moment Correlation

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$$Cov(X, Y) = \mathbf{E}(XY) - \mathbf{E}(X)\mathbf{E}(Y)$$

= $\mathbf{E}((X - \bar{X})(Y - \bar{Y}))$

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Correlation (standardized covariance)

$$r = r_{X,Y} = \frac{Cov(X,Y)}{s_X s_Y}$$
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$$\hat{\beta}_{1} = \frac{\sum_{i=1}^{n} (X_{i} - \bar{X})(Y_{i} - \bar{Y})}{\sum_{i=1}^{n} (X_{i} - \bar{X})^{2}}$$
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If one standardizes *X* and *Y* (i.e. subtract the mean and divide by the standard deviation) before running linear regression, then: $\hat{\beta}_0 = 0$ and $\hat{\beta}_1 = r$ --- *i.e.* $\hat{\beta}_1$ *is the Pearson correlation!*

Comparing Variables

- Linear Regression
- Pearson Product-Moment Correlation
- Multiple Linear Regression
- (Multiple) Logistic Regression
- Ridge Regression (L2 Penalized)
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where $\mathbf{E}(\epsilon_i | X_i) = 0$ and $\mathbf{V}(\epsilon_i | X_i) = \sigma^2$
expected variance

Estimated intercept and slope

$$\hat{r}(x) = \hat{\beta}_0 + \hat{\beta}_1 x \quad \hat{Y}_i = \hat{r}(X_i)$$
Residual: $\hat{\epsilon}_i = Y_i - \hat{Y}_i$

Suppose we have multiple *X* that we'd like to fit to *Y* at once:

$$Y_{i} = \beta_{0} + \beta_{1}X_{i1} + \beta_{2}X_{i2} + \dots + \beta_{m}X_{m1} + \epsilon_{i}$$

If we include and $X_{oi} = 1$ for all i (i.e. adding the intercept to X), then we can say: $Y_i = \sum_{j=0}^m \beta_j X_{ij} + \epsilon_i$

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Suppose we have multiple independent variables that we'd like to fit to our dependent variable: $Y_i = \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \ldots + \beta_m X_{m1} + \epsilon_i$

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To test for significance of
individual coefficient, *j*:

$$t = \frac{\hat{\beta}_{j}}{SE(\hat{\beta}_{j})} = \frac{\hat{\beta}_{j}}{\sqrt{\frac{s^{2}}{\sum_{i=1}^{n} (X_{ij} - \bar{X}_{j})^{2}}}}$$
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Estimating β :
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T-Test for significance of hypothesis:1) Calculate *t*

2) Calculate degrees of freedom:

$$df = N - (m+1)$$

3) Check probability in a *t* distribution:



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3) Check probability in a *t* distribution: (df = v)

Summary

Hypothesis Testing:

A framework for deciding which differences/relationships matter.

- Random Variables
- Distributions
- Hypothesis Testing Framework

Comparing Variables:

Metrics to quantify the difference or relationship between variables.

- Simple Linear Regression, Correlation, Multiple Linear Regression,
- Comparing Variables and Hypothesis Testing
- Regularized Linear Regression (for supervised ML)
- Multiple Hypothesis Testing

Large-Scale Hypothesis Testing

- Findings and Uncertainty
- Hypothesis Testing
- Bonferroni's Cats
- Multi-test Corrections
 - Family-wise Error Rate
 - False-Discovery Rate
- Correlation Metrics
 - Effect Size (coefficient)
 - Significance (whether p-value is below significance level)

Supplement: Not on exam

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> Note: this is a probability here. In simple linear regression we wanted an expectation: r(x) = E(Y|X = x)

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In simple linear regression we wanted an expectation:

$$r(x) = \mathcal{E}(Y|X = x)$$

(i.e. if p > 0.5 we can confidently predict $Y_i = 1$)

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$$logit(p_i) = log\left(\frac{p_i}{1-p_i}\right) = \beta_0 + \sum_{j=1}^m \beta_j x_{ij}$$

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$$logit(p_i) = log\left(\frac{p_i}{1-p_i}\right) = \beta_0 + \sum_{j=1}^m \beta_j x_{ij}$$
$$P(Y_i = 0 \mid X = x)$$
Thus, 0 is class 0
and 1 is class 1.

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(https://www.linkedin.com/pulse/predicting-outcomes-pr obabilities-logistic-regression-konstantinidis/)

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To estimate β , one can use *reweighted least squares:*

(Wasserman, 2005; Li, 2010)

set $\hat{\beta}_0 = \dots = \hat{\beta}_m = 0$ (remember to include an intercept) 1. Calculate p_i and let W be a diagonal matrix where element $(i, i) = p_i(1 - p_i)$. 2. Set $z_i = logit(p_i) + \frac{Y_i - p_i}{p_i(1 - p_i)} = X\hat{\beta} + \frac{Y_i - p_i}{p_i(1 - p_i)}$ 3. Set $\hat{\beta} = (X^T W X)^{-1} X^T W z //$ weighted lin. reg. of Z on Y. 4. Repeat from 1 until $\hat{\beta}$ converges.

Uses of linear and logistic regression

- 1. Testing the relationship between variables given other variables. β is an "effect size" -- a score for the magnitude of the relationship; can be tested for significance.
- 2. Building a predictive model that generalizes to new data. \hat{Y} is an estimate value of Y given X.

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- Building a predictive model that generalizes to new data.

 Ŷ is an estimate value of *Y* given *X*.

 However, unless |*X*| <<< observatations then the model might "overfit".

-> Regularized linear regression (a ML technique)

Statistical Considerations in Big Data

- Correct for multiple tests (Bonferonni's Principle)
- 2. Average multiple models (ensemble techniques)
- 3. Smooth data
- 4. "Plot" data (or figure out a way to look at a lot of it "raw")
- 5. Interact with data

(http://simplystatistics.org/2014/05/22/10-things-statistics-taught-us-about-big-data-analysis/)

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- 6. Know your "real" sample size
- 7. Correlation is not causation
- 8. Define metrics for success (set a baseline)
- 9. Share code and data
- 10. The problem should drive solution

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